

# A Note on the Weight of Multicast Shortest Path Trees

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## Abstract

The weight of a multicast shortest path tree in the complete graph with exponential link weights is expressed as a random variable. The new framework elegantly allows the computation of the average multicast weight.

## 1 Introduction

We consider the problem of computing the cost of multicasting information from a source to  $m$  different nodes in a network containing  $N + 1$  nodes. We assume that the multicast tree rooted at the source, which we label by 0, is the shortest path tree to the  $m$  different nodes. In addition, these  $m$  multicast member nodes are supposed to be uniformly distributed among the  $N$  nodes (different from the source) in the network. Both assumptions are realistic and have been justified earlier [2].

We further confine to a complete graph  $K_{N+1}$  with exponentially distributed link weights with mean 1. The present note is an extension of [3] where the weight  $W_N(m)$  to all  $m = N$  possible multicast members is shown to tend to a Gaussian,

$$\sqrt{N} (W_N(N) - \zeta(2)) \xrightarrow{d} N(0, \sigma_{\text{SPT}}^2),$$

where  $\sigma_{\text{SPT}}^2 = 4\zeta(3)$ . In addition, we recompute the average weight of the multicast tree  $\mathbb{E}[W_N(m)]$  (see eq. (10)) more elegantly than in our previous work [1]. Beside the novel mathematical framework, the work is motivated by the new peer-to-peer protocol HMTP that attaches users to the tree in order to minimize the total cost  $W_N(m)$ . Moreover, in IPTV, an operator is businesswise interested in the cost difference  $W_N(m+k) - W_N(m)$  by adding  $k$  additional customers to a TV channel, given a multicast tree of that TV channel with  $m$  multicast members. The number of hops to the nearest of  $m$  peers is analyzed in [4].

## 2 The weight of the SPT to $m$ nodes

The shortest path tree in the complete graph with exponential link weight is a uniform recursive tree. A uniform recursive tree (URT) of size  $N$  is a random tree rooted at node 0 where at each stage of the growth process a new node is attached uniformly to one of the existing nodes until the total number

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of nodes is equal to  $N$ . The nice properties of the shortest path search process in the complete graph with exponential link weights are explained in [5, Chapter 16]. The path search or discovery process can be described as a Markov process. From the Markov discovery process, the discovery time to the  $k^{\text{th}}$  discovered node from the root equals

$$v_k = \sum_{n=1}^k \tau_n, \quad (1)$$

where the inter-attachment times  $\tau_1, \tau_2, \dots, \tau_k$  are independent, exponentially distributed random variables with parameter  $\lambda_n = n(N+1-n)$  for  $1 \leq n \leq k$ . An arbitrary uniform recursive tree consisting of  $N+1$  nodes and with the root labeled by zero can be represented as

$$(0 \longleftarrow 1)(N_2 \longleftarrow 2) \dots (N_N \longleftarrow N)$$

where  $(N_j \longleftarrow j)$  means that the  $j^{\text{th}}$  discovered node is attached to node  $N_j \in \{0, \dots, j-1\}$ . Hence,  $N_j$  is the predecessor of  $j$  and this relation is indicated by  $\longleftarrow$ . The weight  $W_N(m)$  of the SPT from the root 0 to  $m$  other nodes is with (1) and  $v_0 = 0$  and  $N_1 = 0$ ,

$$W_N(m) = \sum_{j=1}^N (v_j - v_{N_j}) \mathbf{1}_{j \in T_m} = \sum_{j=1}^N \mathbf{1}_{j \in T_m} \sum_{n=N_j+1}^j \tau_n,$$

where  $T_m$  is the subtree of the complete SPT to the  $m$  uniform nodes. In the URT, the integers  $N_j$  for  $1 \leq j \leq N$ , are independent and uniformly distributed over the interval  $\{0, \dots, j-1\}$ . Following the analysis in [3], it is more convenient to use a discrete uniform random variable on  $\{1, \dots, j\}$  which we define as  $A_j = N_j + 1$  such that

$$W_N(m) = \sum_{j=1}^N \mathbf{1}_{j \in T_m} \sum_{n=A_j}^j \tau_n = \sum_{j=1}^N \mathbf{1}_{j \in T_m} \sum_{n=1}^j \mathbf{1}_{\{A_j \leq n\}} \tau_n$$

The set  $\{A_j\}_{1 \leq j \leq N}$  are independent random variables with  $\mathbb{P}[A_j = k] = \frac{1}{j}$  for  $k \in \{1, 2, \dots, j\}$ . If we denote the  $m$  uniformly chosen multicast member nodes by  $U_1, \dots, U_m$ , then node  $j$  is an element of  $T_m$  precisely when there exists an  $i = 1, \dots, m$  such that  $U_i \in \mathcal{T}_j^{(N)}$ , where  $\mathcal{T}_j^{(N)}$  is the subtree of the complete uniform recursive tree rooted at  $j$ . We write  $\vec{U}_m \cap \mathcal{T}_j^{(N)} \neq \emptyset$  for the event that there exists an  $i = 1, \dots, m$  such that  $U_i \in \mathcal{T}_j^{(N)}$ . Then, we have

$$W_N(m) = \sum_{j=1}^N \mathbf{1}_{\{\vec{U}_m \cap \mathcal{T}_j^{(N)} \neq \emptyset\}} \sum_{n=1}^j \mathbf{1}_{\{A_j \leq n\}} \tau_n = \sum_{n=1}^N \tau_n \left( \sum_{j=n}^N \mathbf{1}_{\{\vec{U}_m \cap \mathcal{T}_j^{(N)} \neq \emptyset\}} \mathbf{1}_{\{A_j \leq n\}} \right).$$

The indicator  $\mathbf{1}_{\{\vec{U}_m \cap \mathcal{T}_j^{(N)} \neq \emptyset\}}$  depends on the random variables  $U_1, \dots, U_m$ , which are independent of  $\{A_j\}_{1 \leq j \leq N}$  and of the tree  $\mathcal{T}_j^{(N)}$ . Since the nodes  $1, 2, \dots, j$  are attached before the tree  $\mathcal{T}_j^{(N)}$  starts growing at node  $j$ , the indicator  $\mathbf{1}_{\{\vec{U}_m \cap \mathcal{T}_j^{(N)} \neq \emptyset\}}$  is independent of  $A_1, \dots, A_j$ , but not of  $A_{j+1}, \dots, A_N$ . Of course, the entire recursive tree is independent of the sequence  $\tau_1, \tau_2, \dots, \tau_N$ .

To simplify the notation, we define  $Z_j(m)$  as the number of elements that the vector  $\vec{U}_m$  has in common with the random set  $\mathcal{T}_j^{(N)}$ ,

$$Z_j(m) = |\vec{U}_m \cap \mathcal{T}_j^{(N)}|, \quad j = 1, 2, \dots, N, \quad (2)$$

where we use the notation  $|G|$  for the number of elements of the set  $G$ . We define for  $n \in \{1, \dots, N\}$  the random variables

$$B_n(m) = \sum_{j=n}^N \mathbf{1}_{\{\vec{U}_m \cap \mathcal{T}_j^{(N)} \neq \emptyset\}} \mathbf{1}_{\{A_j \leq n\}} = \sum_{j=n}^N \mathbf{1}_{\{Z_j(m) > 0\}} \mathbf{1}_{\{A_j \leq n\}}, \quad (3)$$

to obtain

$$W_N(m) = \sum_{n=1}^N B_n(m) \tau_n. \quad (4)$$

The  $n$  random variables  $B_1(m), B_2(m), \dots, B_n(m)$  are dependent, but independent from the inter-attachment times  $\{\tau_j\}_{1 \leq j \leq N}$ . Hence, the average weight of the multicast shortest path tree is

$$\mathbb{E}[W_N(m)] = \sum_{n=1}^N \frac{\mathbb{E}[B_n(m)]}{n(N+1-n)}. \quad (5)$$

We can also express the variance (and any other moment)

$$\text{Var}[W_N(m)] = \sum_{n=1}^N \frac{(\mathbb{E}[B_n(m)])^2}{(n(N+1-n))^2} + 2 \sum_{n=1}^N \sum_{l=n}^N \frac{\text{Cov}[B_n(m), B_l(m)]}{n(N+1-n)l(N+1-l)} \quad (6)$$

but, we did not succeed so far in computing  $\text{Cov}[B_n(m), B_l(m)]$ .

### 3 The average weight $\mathbb{E}[W_N(m)]$

The average weight  $\mathbb{E}[W_N(m)]$  requires the computation of  $\mathbb{E}[B_n(m)]$  in which the size of the subtree  $\mathcal{T}_j^{(N)}$  rooted at an arbitrary node  $j$  plays an important role. We first note that

$$\begin{aligned} \mathbb{P}[Z_j(m) > 0] &= 1 - \mathbb{P}[\vec{U}_m \cap \mathcal{T}_j^{(N)} = \emptyset] \\ &= 1 - \mathbb{E}\left[\frac{(N - |\mathcal{T}_j^{(N)}|) \cdots (N + 1 - m - |\mathcal{T}_j^{(N)}|)}{N \cdots (N + 1 - m)}\right] \\ &= 1 - \frac{(N - m)!}{N!} \mathbb{E}\left[(N - |\mathcal{T}_j^{(N)}|) \cdots (N + 1 - m - |\mathcal{T}_j^{(N)}|)\right] \end{aligned} \quad (7)$$

since the event  $\vec{U}_m \cap \mathcal{T}_j^{(N)} = \emptyset$  requires that each of the uniformly chosen multicast member nodes  $U_i$ , for  $i = 1, \dots, m$ , should not lie in  $\mathcal{T}_j^{(N)}$ . Therefore, the mean of the random variable  $B_n(m)$  follows from (3) with  $\mathbb{P}[A_j \leq n] = \frac{n}{j}$  as

$$\begin{aligned} \mathbb{E}[B_n(m)] &= \sum_{j=n}^N \mathbb{E}[\mathbf{1}_{\{A_j \leq n\}}] \mathbb{P}[Z_j(m) > 0] = n \sum_{j=n}^N \frac{\mathbb{P}[Z_j(m) > 0]}{j} \\ &= n \sum_{j=n}^N \frac{1}{j} \left(1 - \frac{(N - m)!}{N!} \mathbb{E}\left[(N - |\mathcal{T}_j^{(N)}|) \cdots (N + 1 - m - |\mathcal{T}_j^{(N)}|)\right]\right). \end{aligned} \quad (8)$$

In [1], we have shown<sup>1</sup> that, for  $j > 0$ ,

$$\mathbb{P}[|\mathcal{T}_j^{(N)}| = n] = \frac{j(N-j)!(N-n)!}{N!(N-j-n+1)!}.$$

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<sup>1</sup>The number of nodes  $N$  and node label  $j$  in [1] should be replaced here by  $N+1$  and by  $j+1$ .

Since  $j$  is never equal to the root 0, it means that the largest subtree tree  $|\mathcal{T}_j^{(N)}|$  is of size  $N$ , but never smaller than  $|\mathcal{T}_j^{(N)}| = 1$  (namely the node  $j$  itself). The probability generating function  $\varphi_{|\mathcal{T}_j^{(N)}|}(z)$  of  $|\mathcal{T}_j^{(N)}|$  is

$$\begin{aligned}\varphi_{|\mathcal{T}_j^{(N)}|}(z) &= \mathbb{E} \left[ z^{|\mathcal{T}_j^{(N)}|} \right] = \sum_{n=1}^N \mathbb{P} \left[ |\mathcal{T}_j^{(N)}| = n \right] z^n = \frac{j(N-j)!}{N!} \sum_{n=1}^N \frac{(N-n)!}{(N-j-n+1)!} z^n \\ &= \frac{j(N-j)!}{N!} \sum_{k=0}^{N-1} \frac{k!}{(k-(j-1))!} z^{N-k}\end{aligned}$$

Thus,

$$\begin{aligned}\mathbb{E} \left[ (N - |\mathcal{T}_j^{(N)}|) \cdots (N + 1 - m - |\mathcal{T}_j^{(N)}|) \right] &= \frac{d^m}{dz^m} E \left[ z^{N-|\mathcal{T}_j^{(N)}|} \right] \Big|_{z=1} = \frac{d^m}{dz^m} E \left[ z^N \varphi_{|\mathcal{T}_j^{(N)}|}(z^{-1}) \right] \Big|_{z=1} \\ &= \frac{j(N-j)!}{N!} \sum_{k=0}^{N-1} \frac{k!}{(k-(j-1))!} \frac{d^m}{dz^m} E \left[ z^k \right] \Big|_{z=1} \\ &= \frac{j(N-j)!}{N!} \sum_{k=0}^{N-1} \frac{k!}{(k-(j-1))!} \frac{k!}{(k-m)!}\end{aligned}$$

Substituted in (8) gives

$$\mathbb{E}[B_n(m)] = \sum_{j=n}^N \frac{n}{j} \left( 1 - \frac{(N-m)!j(N-j)!}{(N!)^2} \sum_{k=0}^{N-1} \frac{(k!)^2}{(k-(j-1))!(k-m)!} \right)$$

and in (5)

$$\mathbb{E}[W_N(m)] = \sum_{n=1}^N \frac{1}{N+1-n} \sum_{j=n}^N \frac{1}{j} \left( 1 - \frac{(N-m)!j(N-j)!}{(N!)^2} \sum_{k=0}^{N-1} \frac{(k!)^2}{(k-(j-1))!(k-m)!} \right)$$

The first sum is, after reversal of the summations and using the identity  $\sum_{j=1}^N \frac{1}{j} \sum_{k=N+1-j}^N \frac{1}{k} = \sum_{k=1}^N \frac{1}{k^2}$  proved [3, Appendix], equal to

$$\sum_{n=1}^N \frac{1}{(N+1-n)} \sum_{j=n}^N \frac{1}{j} = \sum_{k=1}^N \frac{1}{k^2}$$

The second sum is

$$\begin{aligned}Y_m &= \frac{(N-m)!}{(N!)^2} \sum_{n=1}^N \frac{1}{N+1-n} \sum_{j=n}^N (N-j)! \sum_{k=0}^{N-1} \frac{(k!)^2}{(k-(j-1))!(k-m)!} \\ &= \frac{(N-m)!}{(N!)^2} \sum_{n=1}^N \frac{1}{N+1-n} \sum_{k=0}^{N-1} \frac{(k!)^2}{(k-m)!} \sum_{j=n}^N \frac{(N-j)!}{(k-(j-1))!}\end{aligned}$$

Application of the identity

$$\sum_{j=n}^m \frac{(a-j)!}{(b-j)!} = \frac{1}{a+1-b} \left\{ \frac{(a-n+1)!}{(b-n)!} - \frac{(a-m)!}{(b-m-1)!} \right\} \quad (9)$$

gives

$$\begin{aligned} Y_m &= \frac{(N-m)!}{(N!)^2} \sum_{n=1}^N \sum_{k=0}^{N-1} \frac{(k!)^2 (N-n)!(N-k-1)!}{(k-m)!(k+1-n)!(N-k)!} \\ &= \frac{(N-m)!}{(N!)^2} \sum_{k=0}^{N-1} \frac{(k!)^2}{(k-m)!(N-k)} \sum_{n=1}^N \frac{(N-n)!}{(k+1-n)!} \end{aligned}$$

Again, using (9) yields

$$\begin{aligned} Y_m &= \frac{(N-m)!}{N!} \sum_{k=0}^{N-1} \frac{k!}{(k-m)!(N-k)^2} \\ &= \frac{(N-m)!}{N!} \sum_{k=1}^N \frac{(N-k)!}{(N-m-k)!k^2} = \frac{1}{\binom{N}{m}} \sum_{k=1}^N \binom{N-k}{m} \frac{1}{k^2} \end{aligned}$$

Hence, we arrive at

$$\mathbb{E}[W_N(m)] = \sum_{k=1}^N \frac{1}{k^2} - Y = \sum_{k=1}^N \frac{\binom{N}{m} - \binom{N-k}{m}}{\binom{N}{m}} \frac{1}{k^2} \quad (10)$$

$$= \sum_{j=1}^m \frac{1}{N+1-j} \sum_{k=j}^N \frac{1}{k} \quad (11)$$

where the last formula (11) was found in [1]. Equality of both formulae is proved in Appendix A.

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## References

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## A Proof that (10) equals (11)

Writing the difference in  $Y$  as  $\Delta Y = Y_m - Y_{m-1} = -\Delta W$ ,

$$\Delta Y = \frac{1}{N!} \sum_{k=1}^N \left[ \frac{(N-m)!}{(N-m-k)!} - \frac{(N-m+1)!}{(N-m+1-k)!} \right] \frac{(N-k)!}{k^2}$$

and with

$$B = \frac{(N-m)!}{(N-m-k)!} - \frac{(N-m+1)!}{(N-m+1-k)!} = \frac{-(N-m)!k}{(N-m+1-k)!}$$

we see that

$$\begin{aligned} \Delta Y &= \frac{-(N-m)!}{N!} \sum_{k=1}^N \frac{(N-k)!}{(N-m+1-k)!} \frac{1}{k} \\ &= \frac{-(N-m)!}{N!} \sum_{k=1}^{N-m+1} \frac{(N-k)!}{(N-m+1-k)!} \frac{1}{k} \end{aligned}$$

Using the identity

$$\sum_{j=n}^b \frac{1}{j} \frac{(a-j)!}{(b-j)!} = \frac{a!}{b!} \sum_{j=n}^b \frac{1}{j} - \frac{a!}{(b-n)!} \sum_{q=0}^{a-b-1} \frac{1}{a-b-q} \frac{(a-q-n)!}{(a-q)!}$$

which is valid for all integers  $a, b$  and  $n$ , gives

$$\sum_{j=1}^{N-m+1} \frac{1}{j} \frac{(N-j)!}{(N-m+1-j)!} = \frac{N!}{(N-m+1)!} \sum_{j=1}^{N-m+1} \frac{1}{j} - \frac{N!}{(N-m)!} \sum_{q=0}^{m-2} \frac{1}{m-1-q} \frac{1}{(N-q)}$$

Since

$$\begin{aligned} \sum_{q=0}^{m-2} \frac{1}{m-1-q} \frac{1}{(N-q)} &= \frac{1}{N-m+1} \sum_{q=0}^{m-2} \frac{1}{m-1-q} - \frac{1}{N-m+1} \sum_{q=0}^{m-2} \frac{1}{(N-q)} \\ &= \frac{1}{N-m+1} \sum_{j=1}^{m-1} \frac{1}{j} - \frac{1}{N-m+1} \sum_{j=N-m+2}^N \frac{1}{j} \end{aligned}$$

we have

$$\begin{aligned} \Delta Y &= -\frac{1}{(N-m+1)} \sum_{j=1}^{N-m+1} \frac{1}{j} + \frac{1}{N-m+1} \sum_{j=1}^{m-1} \frac{1}{j} - \frac{1}{N-m+1} \sum_{j=N-m+2}^N \frac{1}{j} \\ &= -\frac{1}{N-m+1} \sum_{j=m}^N \frac{1}{j} \end{aligned}$$

From the original expression (11), we immediately find that  $\Delta W = \frac{1}{N+1-m} \sum_{k=m}^N \frac{1}{k}$  which proves equality in the differences since  $\Delta Y = -\Delta W$ . Equality of (11) and (10) then follows since for  $m = N$ , both expressions are equal.  $\square$