1	Individualized epidemic spreading models predict epilepsy surgery				
2	outcomes: a pseudo-prospective study				
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<sup>15</sup> Personalized Medicine, Seizure Modeling, Whole-brain Modeling

# ABSTRACT

<sup>16</sup> Epilepsy surgery is the treatment of choice for drug-resistant epilepsy patients, but up to 50% of patients

<sup>17</sup> continue to have seizures one year after the resection. In order to aid presurgical planning and predict

<sup>18</sup> postsurgical outcome on a patient-by-patient basis, we developed a framework of individualized

<sup>19</sup> computational models that combines epidemic spreading with patient-specific connectivity and

<sup>20</sup> epileptogeneity maps: the Epidemic Spreading Seizure and Epilepsy Surgery framework (ESSES).

ESSES parameters were fitted in a retrospective study (N = 15) to reproduce invasive

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electroencephalography (iEEG)-recorded seizures. ESSES reproduced the iEEG-recorded seizures, and 22 significantly better so for patients with good (seizure-free, SF) than bad (non-seizure-free, NSF) 23 outcome. We illustrate here the clinical applicability of ESSES with a *pseudo-prospective study* 24 (N = 34) with a blind setting (to the resection strategy and surgical outcome) that emulated presurgical 25 conditions. By setting the model parameters in the retrospective study, ESSES could be applied also to 26 patients without iEEG data. ESSES could predict the chances of good outcome after any resection by 27 finding patient-specific model-based optimal resection strategies, which we found to be smaller for SF 28 than NSF patients, suggesting an intrinsic difference in the network organization or presurgical 29 evaluation results of NSF patients. The actual surgical plan overlapped more with the model-based 30 optimal resection, and had a larger effect in decreasing modeled seizure propagation, for SF patients than 31 for NSF patients. Overall, ESSES could correctly predict 75% of NSF and 80.8% of SF cases 32 pseudo-prospectively. Our results show that individualised computational models may inform surgical 33 planning by suggesting alternative resections and providing information on the likelihood of a good 34 outcome after a proposed resection. This is the first time that such a model is validated with a fully 35 independent cohort and without the need for iEEG recordings. 36

## **AUTHOR SUMMARY**

Individualized computational models of epilepsy surgery capture some of the key aspects of seizure 37 propagation and the resective surgery. It is to be established whether this information can be integrated 38 during the presurgical evaluation of the patient to improve surgical planning and the changes of a good 39 surgical outcome. Here we address this question with a pseudo-prospective study that applies a 40 computational framework of seizure propagation and epilepsy surgery - the ESSES framework- in a 41 pseudo-prospective study mimicking the presurgical conditions. We found that, within this 42 pseudo-prospective setting, ESSES could correctly predict 75% of NSF and 80.8% of SF cases. This 43 finding suggests the potential of individualised computational models to inform surgical planning by 44 suggesting alternative resections and providing information on the likelihood of a good outcome after a 45 proposed resection. 46

# INTRODUCTION

Surgical resection is often the most effective treatment to achieve seizure control for patients with drug-resistant focal epilepsy. The surgery requires the generation of an hypothesis of the epileptogenic 48 zone (EZ) by means of extensive presurgical evaluations, and its subsequent removal or disconnection 49 during surgery (Lüders, Najm, Nair, Widdess-Walsh, and Bingman (2006)). Despite extensive 50 investigations, there has only been a slight improvement in prognosis over the past two decades 51 (Baxendale et al. (2019); Jehi et al. (2015)), and between 30 to 50% of the patients who undergo surgery 52 continue to have seizures one year later, depending on etiology and location of the EZ (Englot et al. 53 (2015)). A key conceptual change in recent years is the notion of *epileptogenic networks*, which takes 54 into account the complex interplay between different brain regions in promoting and inhibiting seizure 55 generation and propagation (Bartolomei et al. (2017); Kramer and Cash (2012); van Diessen, Diederen, 56 Braun, Jansen, and Stam (2013)). As a consequence, the effect of a given surgery is to be measured 57 against the whole epileptogenic network: a small resection involving heavily connected regions may have 58 widespread effects, but it may also be compensated for by the remaining network (Hebbink, Meijer, 59 Huiskamp, van Gils, and Leijten (2017); Nissen et al. (2018)). This perspective aligns with the 60 commonly accepted view that large-scale brain organization can be regarded as an emerging 61 phenomenon taking place on a complex network, which has spurred numerous data- and model-based 62 studies (Seguin, Jedynak, et al. (2023); Seguin, Sporns, and Zalesky (2023)). Several network-based 63 studies have found group-level differences between seizure-free and non-seizure-free patients (da Silva et 64 al. (2020); Nissen et al. (2018); Taylor et al. (2018)), with removal of pathological hub (i.e. central) 65 regions typically associated with seizure-freedom (Nissen et al. (2017)). These results highlight the need 66 to consider patient-specific connectivity (van den Heuvel and Sporns (2019)) in order to tailor the surgery 67 to each patient (Gerster et al. (2021)). 68

A data-driven manner to study the relation between individual brain networks and surgical outcomes
involves *computational models of seizure dynamics*, which allow us to simulate seizure propagation *in silico*. Different resection strategies can be tested on the computational model before the actual surgery
(Goodfellow et al. (2016); Hutchings et al. (2015); V. Jirsa et al. (2017); Laiou et al. (2019); Lopes et al.

<sup>73</sup> (2017); Nissen et al. (2021); Olmi, Petkoski, Guye, Bartolomei, and Jirsa (2019); Proix, Bartolomei,

<sup>74</sup> Chauvel, Bernard, and Jirsa (2014); Sinha et al. (2017); Taylor, Kaiser, and Dauwels (2014)). The models

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can be fitted to patient-specific data of brain structure and seizure dynamics, allowing us to tailor the 75 resection strategy for each patient. Within this perspective, previous studies have obtained remarkable 76 success at a group level: Sinha et al. (2017) found that the removal of regions identified as epileptogenic 77 according to an EEG-brain network dynamical model predicted surgical outcome with 81.3% accuracy. 78 Proix, Bartolomei, Guye, and Jirsa (2017), using a seizure model known as the epileptor (V. K. Jirsa, 79 Stacey, Quilichini, Ivanov, and Bernard (2014)) based on MRI (magnetic resonance imaging) 80 connectivity, found significant differences in the overlap between the model-based propagation zone and 81 the area sampled by iEEG between patients with good (Engel class I) and bad (Engel class III) outcomes 82 at the group level. Subsequent studies also found a better match between the modeled and clinically 83 observed epileptogenic regions for seizure-free than non-seizure-free patients (Makhalova et al. (2022); 84 Vattikonda et al. (2021)). Similarly, Sip et al. (2021) simulated patient-specific resection strategies by 85 means of virtual resections, and found that virtual resections in their model correlated with surgical 86 outcome, with larger effects found for patients with good outcome (Engel classes I and II). In an 87 independent study, Goodfellow et al. (2016) also found significant differences in the model prediction 88 between Engel class I and class IV patients, using an electrocorticogram modeling framework. 89

Following the same rationale, we developed a computational model of seizure propagation and epilepsy 90 surgery based on epidemic spreading dynamics and patient-specific MEG brain connectivity (Millán et 91 al. (2022)), to which we refer here as the Epidemic Spreading Seizure and Epilepsy Surgery model 92 *ESSES*). Epidemic models describe the spread of an infectious agent through a network. Epidemic 93 processes on fixed networks have a rich mathematical history (Pastor-Satorras, Castellano, Van Mieghem, 94 and Vespignani (2015)) with a plethora of models that can be exploited for epilepsy surgery optimization 95 (Millán et al. (2022); Nissen et al. (2021)). Although such models ignore the underlying bio-physical 96 processes that lead to seizure generation and propagation, they describe the basic rules that govern 97 spreading processes. In previous studies (Millán et al. (2022, 2023)), we found that epidemic spreading 98 models could reproduce stereotypical patterns of seizure propagation as recorded via invasive 99 electroencephalography (iEEG) recordings. Moreover, once fitted with patient-specific data, ESSES 100 could identify alternative resection strategies, either of smaller size or at a different location than the 101 actual surgery (Millán et al. (2022); Nissen et al. (2021)). In a more recent study Millán et al. (2023), we 102 showed that the goodness-of-fit of ESSES seizures to those recorded via iEEG predicted surgical 103

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outcome –with an area under the curve of 88.6% – indicating that ESSES not only reproduces the basic 104 aspects of seizure propagation, but it also captures the differences, either in the location of the resection 105 area relative to the EZ, or intrinsically in the iEEG or MEG data, between patients with good and bad 106 outcome. Importantly, ESSES's global parameters were defined at the population level, and the model 107 was individualized for each patient via patient-specific MEG networks, which characterized the local 108 spreading probabilities. As a consequence, ESSES can be extended to patients without iEEG recordings, 109 in contrast to previous modeling studies, which typically required the existence of patient-specific iEEG 110 data to individualize the model for each patient (Bernabei et al. (2023); Gunnarsdottir et al. (2022); 111 Makhalova et al. (2022); Proix et al. (2017); Runfola, Sheheitli, Bartolomei, Wang, and Jirsa (2023); 112 Sinha et al. (2017); Y. Wang et al. (2023)). IEEG allows for a highly resolved description of seizure 113

<sup>114</sup> dynamics, but its spatial sampling is sparse and it is highly invasive. Consequently, it is only part of the <sup>115</sup> presurgical evaluation in a selection of patients.

Here we performed a pseudo-prospective blind study (34-patient validation cohort) to validate the clinical 116 applicability of ESSES to a) identify model-based optimal resection strategies and b) predict the 117 likelihood of a good outcome after a proposed resection strategy, on a patient-by-patient basis. In order to 118 emulate the clinical presurgical conditions, the research team was blind to the patients' postsurgical data, 119 namely the resection area and surgical outcome, during ESSES's analyses, and the multimodal 120 presurgical information available for each patient was integrated into ESSES. ESSES can identify 121 resection strategies that perform optimally in the model, i.e. by minimizing modeled seizure propagation, 122 for a given resection size. We refer to these resections as *optimal resections*, in agreement with previous 123 works (An, Bartolomei, Guye, and Jirsa (2019); Millán et al. (2022); Nissen et al. (2021); Sinha et al. 124 (2017)). ESSES can also simulate the effect of a given resection *in silico*. Within this set-up, we tested 125 three hypotheses: a) seizure-free (SF) patients would have smaller optimal resections than 126 non-seizure-free (NSF) patients, b) SF patients would have a larger overlap between optimal and planned 127 (clinical) resections, and c) the planned resection would have a larger effect (in ESSES) for SF than for 128 NSF patients. We found that these three ESSES biomarkers, namely the size of the optimal resection, 129 their overlap with the planned resection, and the effect of the planned resection on ESSES seizures, 130 provided estimates of the likelihood of a good outcome after the surgery, as well as suggesting alternative 131 resection strategies that performed optimally in the model. We envisage that the implementation of a 132

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modeling scheme such as ESSES in clinical practice may inform the planning of epilepsy surgery. 133

Different surgical plans can be tested with ESSES for each patient, such that strategies that lead to a large 134 decrease of propagation in the model are more likely to lead to seizure freedom. ESSES may also suggest 135 optimal (alternative) resection strategies, for cases where ESSES predicts a bad outcome with the planned 136 resection. Optimal strategies can then lead to new surgical plans, the effect of which can then be tested in 137

ESSES again. 138

# RESULTS

Here we validated the clinical applicability of ESSES to A) identify optimal resection strategies that may 139 improve surgical outcomes and **B**) provide estimates of the probability of postsurgical seizure freedom, 140 given a surgical plan. The key goal of ESSES is to identify surgical candidates who would have a bad 141 outcome (NSF patients) so that the surgical plan can be adjusted. This study combined a retrospective 142 analysis on a modeling cohort (N = 15) that was used to set the model hyperparameters (following our 143 previous retrospective study (Millán et al. (2023)) on this same cohort), and a pseudo-prospective study 144 on a validation cohort (N = 34) to validate ESSES findings and to emulate its clinical application in a 145 blind set-up that mimics the clinical presurgical conditions. The researchers were blind to the performed 146 surgery and surgical outcome during the application of ESSES to the validation cohort. 147

The study was performed as follows: 148

1. Seizure model: definition and fitting (modeling cohort). An SIR-type of epidemic spreading 149 process modeled seizure propagation over patient-specific brain connectivity. IEEG data from the 150 modeling cohort was used to fit the global parameters of the spreading model so that 151 ESSES-modeled seizures matched those recorded via iEEG, as shown in figure 1A. 152 2. Individualized ESSES framework: patient-specific models. ESSES was individualized for each 153

patient: patient-specific MEG brain connectivity defined the network on which ESSES computed 154 seizure propagation. Multi-modal patient-specific data, available from presurgical evaluations, 155 defined the seed regions (i.e. the seizure onset regions) based on *epileptogenicity* or *seed-probability* 156 maps. 157

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Alternative resection strategies (aim A). ESSES incorporates an optimization algorithm to
 determine model-based optimal resection strategies for each patient. These acted as a benchmark
 against which the planned resection for each patient could be tested. These resections were optimal
 in the model in the sense that they minimized modeled seizure propagation.

4. Simulation of the planned resection strategy (aim B). The resection plan for each patient was
 simulated in ESSES with a virtual resection that emulated the actual surgical resection, and the
 subsequent decrease in seizure propagation was measured.

5. Statistical analyses (aim B). We compared ESSES's predictions (steps 3 and 4) between patients
 with good and bad outcome. We defined the NSF class as the positive class for classification and
 prediction testing.

This analysis pipeline was first implemented in the modeling cohort in a retrospective study that served to set all model hyperparameters. Then, steps 2 - 5 were applied to the validation cohort in a pseudo-prospective study with a blind set-up. The pipeline for the model implementation, detailing at which step the de-blinding of each data-type took place, is illustrated in figure 2. A detailed pipeline including also the model set-up (modeling cohort) is also included as Supp. figure 7.

# 173 Seizure propagation as an epidemic spreading process

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We modeled seizure propagation by a Susceptible-Infected-Recovered (SIR) epidemic process, as 192 illustrated in figure 1. The S-I-R states account respectively for the healthy (pre-ictal), ictal and healthy 193 (post-ictal) states, coupled with patient-specific brain connectivity (derived from MEG data) to define the 194 local spreading probabilities. The SIR model describes the spread of an infection from an initial set of 195 infected nodes, the seed regions, to the other nodes in the network, and the recovery of the infected 196 nodes, without re-infections (Barrat, Barthelemy, and Vespignani (2008); Pastor-Satorras et al. (2015)). 197 Here we confined ourselves to one of the simplest compartmental SIR models, using a discrete-time 198 setting where the spreading probability from node i to node j corresponded to the coupling strength  $w_{ij}$ 199 on the patient-specific brain network and where the recovery probability  $\gamma$  was set to be equal for all 200 nodes. The brain network was initially thresholded (by setting the weakest links to zero) at different 201 densities  $\rho$  indicating the fraction of non-zero links remaining in the network after thresholding (see 202 Methods section and Supp. section 5). 203

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A Sketch of ESSES's parameter-fitting scheme. The parameters controlling seizure propagation, namely the Figure 1. 174 density of links in the network  $\rho$  and the global recovery probability  $\gamma$ , were set so as to maximize the similarity between 175 ESSES-modeled seizures and iEEG-recorded ones for the modeling cohort (eq. 1). Seizures were simulated via SIR dynamics 176 over MEG patient-specific brain networks, and setting the resection area as the seed of epidemic spreading. **B**  $\bar{C}(\rho, \gamma)$  map 177 displaying the average model fit (modeling cohort). The data points indicate the parameters corresponding to the best individual 178 fit for each patient, with circles (triangles) indicating SF (NSF) cases (corresponding C values can be seen in Supp. figure 2). 179 Most individual best fits (data-points) fall within the same region (SIR phase transition) but there is large variability (in fact, we 180 found low signal to noise ratios of approx. 1/2, see Supp. figure 3A). The blue square marks the maximum of the goodness-of-181 fit, and the corresponding  $(\rho, \gamma)$  values were used for the subsequent analyses. The y-axis is shown using a logarithmic scale. 182

The two control parameters of ESSES are thus the global recovery probability  $\gamma$  and the network density  $\rho$ . We followed the inference method presented in our previous study (Millán et al. (2023)) to fit the model parameters to iEEG-recorded seizures of the modeling cohort. We note that the modeling framework as presented here differs slightly from the one in Millán et al. (2023), which included an extra parameter to set the global spreading rate. The details of the model fit can be found in the methods == D R A F T

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Figure 2. Processing and analysis pipeline. The patient data were processed in three different steps (blue boxes) for the 183 validation cohort. Firstly, ESSES's key ingredients, the patient-specific MEG brain network and the seed-likelihood map, were 184 processed. The research team remained blind to the resection area and outcome of each patient. The first analysis (AIM A: 185 Optimization of alternative resections, pink boxes) then took place and the first result (Result 1: Size of the optimal resection 186  $R_{op}$ ) was obtained. Then, the patients's resection areas were processed (de-blinding step 1) and the second result was obtained 187 (Result 2: overlap of  $R_{op}$  with the resection area, RA). AIM B (Simulation of the resection plan, yellow boxes) could then take 188 place: the simulation of the resection plan, by performing a virtual resection of the resection area. The third and final result 189 (Result 3: Decrease of spreading  $\delta IR(RA)$ ) was then obtained. Then, the second and final de-blinding took place to recover 190 the outcome of each patient and perform the statistical analyses. 191

section, and the fit results are reported in the supplementary information (Supp. section 5.2, see also
Supp. figures 2 and 3).

The degree of similarity between the ESSES and iEEG seizures was measured with the *goodness-of-fit* 

 $_{212}$   $C(\rho, \gamma)$  (eq. 1). The resulting diagram resembled a familiar phase transition (figure 1B), with an interface

of high goodness-of-fit (yellow regions) corresponding to a roughly constant spreading-to-recovery ratio

 $\rho/\gamma = \text{const}$ , in agreement with other studies (Moosavi, Jirsa, and Truccolo (2022)). The maximum

<sup>215</sup> goodness-of-fit is indicated by a blue square in figure 1B, and sets the working point of ESSES for the

remaining analyses. At this working point, the SF group presented a significantly better fit than the NSF group (p = 0.04, see Supp. table 3 and Supp. figure 3B for details).

A ROC classification analysis indicated a good classification (AUC = 0.79, see Supp. table 4 and Supp. 218 figure 3C) between the SF and NSF groups. At the optimal classification point (Youden criterion, Supp. 219 figure 3D), all NSF patients were correctly identified. The high sensitivity suggests that all patients 220 identified as SF by ESSES could proceed to surgery with high expectations (100% in this group) of a 221 good outcome. On the contrary, patients identified as NSF should be examined further (e.g. by 222 performing further presurgical evaluations or considering other resection plans) as they had a 57% chance 223 of bad outcome with the proposed surgery (to be compared with a 26% chance of bad outcome expected 224 simply from the relative group sizes). 225

# 228 Presurgical hypothesis of the seed regions

A key ingredient of ESSES is the definition of the epileptogenic or seed regions. Here we defined 229 epileptogenicity or seed-probability maps  $SP_i$ , indicating the probability that each brain region i gave 230 rise to a seizure. The *seed-probability* maps integrated patient-specific multimodal presurgical 231 information (encoded in the local patient database (Castor Electronic Data Capture. (n.d.))) in a 232 quantitative and systematic manner that was adapted for each patient to include the data from the 233 presurgical evaluations that they had undergone (see Methods section and Supp. section 4 for details). 234 The resulting seed-probability maps for two representative patients (modeling cohort) are illustrated in 235 figure 3B,D together with the corresponding resection areas (panels A, C). The seed-probability maps 236 show wider spatial patterns than the resection areas, and may involve several lobes in both hemispheres. 237 The resection areas for the two cases shown here were contained within the most likely seed regions. In 238 general, the resection areas had a larger seed-probability than expected by chance for all patients. We did 239 not find significant differences in the overlap between the resection areas and the seed-probability maps 240 between SF and NSF patients (see Supp. figure 1). 241

## 242 Optimal resection strategies

ESSES can derive individualized alternative resection strategies –that minimize modeled seizure

<sup>254</sup> propagation– via an optimization algorithm based on simulated annealing (Millán et al. (2022); Nissen et

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Figure 3. Seed-probability maps. Resection areas (left) and seed-probability maps (right) as derived from the database with presurgical information for two representative cases from the modeling cohort: patient 3 (SF, top) and 6 (NSF, bottom).

al. (2021)). The optimization algorithm parameters were set on the modeling cohort data (see Methods
 section for the algorithm details, and Supp. section 5.3 and Supp. figures 4 and 5 for the modeling cohort
 results), and the algorithm was then applied to the validation cohort in a blind setting.

The optimization algorithm searched for resections R of increasing size S(R) that minimized the *seed* efficiency  $E_R(\text{seed})$ , i.e. the average distance (on the network) from the seed nodes to the other network nodes. This procedure exploits the link between epidemic spreading dynamics and network structure, such that spreading to a region is strongly influenced by its distance to the seed (Pastor-Satorras et al. (2015)). In figure 4A we show the normalized seed efficiency  $e_R(\text{seed})$ , which is normalized to the seed efficiency in the unresected network so as to diminish differences due to seed extent and initial efficiency.  $e_R(\text{seed})$  decreased with the size of the resection for all patients. At the group level, the SF group showed

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Figure 4. Optimal (alternative) resection strategies (validation cohort). Effect of optimal virtual resections of size S(R) as 243 measured by A the normalized seed efficiency  $e_R$  (seed), and B normalized decrease in seizure propagation  $\delta IR(R)$ . Blue 244 dashed lines stand for NSF patients, and pink solid lines for SF patients. Thin lines show individual patients, and darker wide 245 lines the group averages, with shaded areas indicating the standard deviations. The apparent darker pink line at the top of the 246 plot arises from overlap of several individual lines. C-H Group level comparison of the size of optimal resections  $S(R_{op})$  (C-E) 247 and their overlap with the resection area  $Ov(R_{op}, RA)$  (F-H). Panels C and F show the distribution of values of each patient 248 group, with significance results obtained with exact two-sided Wilcoxon ranksum tests. Panels D and G show the corresponding 249 ROC classification analyses, where TPR and FPR stand respectively for the true positive (NSF cases classified as NSF) and 250 false positive (SF cases classified as NSF) rates. Finally, panels E and H show the confusion matrices corresponding to the 251 optimal point (Youden criterion, black asterisks in the middle panels) of the ROC curves. 252

a significantly smaller  $e_R$  (seed) than the NSF group (repeated measures ANOVA test, F(19) = 37.95,  $p < 10^{-89}$ ), for all considered seed sizes except S(R) = 1. Moreover, the effect of increasing the resection size on  $e_R$  (seed) was larger for the SF than for the NSF group (F(19) = 3.78,  $p < 10^{-6}$ ).

The actual effect of a resection R on modeled seizure propagation was quantified by measuring the 268 *normalized decrease in seizure propagation* due to the resection,  $\delta IR(R)$  (figure 4B), again relative to 269 propagation on the unresected network. Seizure propagation depended heavily on the seed realization 270 such that a bi-stable regime emerged in which ESSES seizures either propagated macroscopically or died 271 locally (an exemplary case is shown in Supp. figure 4). Thus, results reported here were averaged over 272 300 independent realizations of the seed regions and SIR dynamics. At the group level, the SF group 273 presented a larger decrease in seizure propagation (F(19) = 25.88,  $p < 10^{-65}$ ), and a larger effect of 274 increasing the resection size  $(F(19) = 2.90, p = 4 \cdot 10^{-5})$ . There were large differences in the 275 dependence of  $\delta IR(R)$  on the resection size between different patients. Whereas in the majority of the 276 cases  $\delta IR(R)$  increased roughly exponentially with S(R), for several patients there was an abrupt 277 (discontinuous) jump at a given resection size. 278

We defined the *optimal resection*  $R_{op}$  as the one leading to a 90% decrease in seizure propagation, 279  $\delta IR(R_{op}) = 0.90$ . The SF group had significantly smaller optimal resections, and these presented a 280 significantly larger overlap with the actual resection strategy  $Ov(R_{op}, RA)$  (see panels C and F of figure 281 4, and table 1), than the NSF group. We found good classification results using either of these variables to 282 classify between the SF and NSF groups (AUC = 0.71, 0.69 respectively for  $S(R_{op})$  and  $Ov(R_{op}, RA)$ , 283 see figure 4D,G). Both variables led to very similar classification results at the optimal classification 284 point (Youden criterion), correctly identifying 6/8 NSF cases (panels E and H). The classification results 285 for the validation cohort are summarized in table 2 (see Supp. table 4 for the modeling cohort results). 286

In summary, these results indicate that the planned resection strategy (accounted for here by the resection area) presented a larger overlap with the optimal resection for patients with good outcome. In particular, 90.0% of SF and 42.9% of NSF patients were correctly classified by  $Ov(RA, R_{op})$ . Remarkably, ESSES could also distinguish between SF and NSF patients without taking into account the information of the surgical plan. In fact, up to 90.4% of SF and 46% of NSF patients were correctly identified by  $S(R_{op})$  (in relation to only a 76.5% SF-chance and 23.5% NSF-chance according to the group ratios). As this analysis did not depend on the planned resection strategy, a bad prognosis would be indicative of the need == D R A F T

to perform a more exhaustive presurgical evaluation, and potentially imply an unavoidable
 non-seizure-free outcome after any surgery.

<sup>296</sup> Finally, we note that almost equivalent results may be obtained by considering the disconnecting
<sup>297</sup> resection, i.e. the smallest resection leading to disconnection of the seed, instead of the optimal resection
<sup>298</sup> (see Supp. section 5.4 and Supp. figure 6). This is due to the strong link between network topology and
<sup>299</sup> emergent SIR dynamics, a result that can be used to speed up computations considerably, by using a
<sup>300</sup> purely network-based analysis of the effect of different resection strategies.

# 301 Simulation of the surgical plan

We simulated the effect of the planned surgery in ESSES for each patient by performing virtual 302 resections of the resection area, which was considered as a proxy for the surgical plan here. We report 303 here on the results for the validation cohort (figure 5), results for the modeling cohort can be found in the 304 supplementary information (Supp. section 7, Supp. figure 8). As in previous sections, all modeling 305 details had already been set during the modeling step. The effect of the resection strategy on (modeled) 306 seizure propagation,  $\delta IR(RA)$ , was significantly larger for the SF than the NSF group (figure 5B, table 307 1). A ROC classification analysis revealed a good classification between the two groups (AUC = 0.78, 308 figure 5C) and at the optimal point (Youden criterion, black asterisk in panel C) the majority of the 309 patients were correctly identified (figure 5D, table 2). In particular, there was a 91.3% chance that a 310 patient classified as SF had a good outcome, and a 54.5% chance that a patient classified as NSF had a 311 bad outcome, compared to a 76.5% and 23.5% chance based on the relative group sizes. 312

# 321 Prediction of surgical outcome

The classification analyses in the previous sections were informed by each patient's surgical outcome. In a prospective setting the outcome for the patient is not yet known, and thus cannot be used to build the classification model. In order to emulate a true prospective setting, we performed a *prediction analysis* based on leave-one-out crossvalidation. That is, in order to predict the outcome of each patient of the validation cohort, a prediction model was built using data from the remaining 33 cases. Results from this analysis are shown in figure 6, with the statistical details reported in table 3. The prediction results were slightly worse than the classification ones (previous sections), particularly for the NSF class where there

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Figure 5. Simulation of the planned resection strategy (validation cohort). A The top panel shows seizure propagation IR313 before (left point cloud for each patient) and after (right point clouds) the resection, for 300 iterations of the seed regions, for 314 each patient. The bottom panel shows the average relative decrease in seizure propagation  $\delta IR(RA)$ , with errorbars given by 315 the standard deviation over seed iterations. **B** Comparison of the relative decrease in seizure propagation  $\delta IR(RA)$  between 316 the SF and NSF groups. Each point corresponds to one patient. C ROC curve of the group classification based on  $\delta IR(RA)$ . 317 TPR and FPR indicate respectively the true positive (NSF cases classified as NSF) and false positive (SF cases classified as 318 NSF) rates. D Classification results for the optimal point (black asterisk in panel C) of the ROC curve according to the Youden 319 criterion. 320

was a 12.5% reduction in the group size. In any case, respectively 4, 5 and 5 NSF cases and 19, 18 and 21
SF cases were correctly identified by each ESSES biomarker (figure 6A). Moreover, 75% of NSF cases
(6/8) and only 19.2% (5/26) of SF cases were identified by two or more biomarkers as NSF (figure 6B).
For this cohort, if ESSES predicted a good outcome with at least two markers, there was a 80.8% chance

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of seizure freedom after the surgery (compared to a 76.5% expectancy of surgery success according to the 333 group rates). Conversely, if the model predicted a bad outcome, then there was a 75% chance that the 334 surgery would fail (compared to a 23.5% expectancy of surgery failure according to the group rates). In 335 clinical practice, a good ESSES prediction could then be interpreted as a large (80.8%) chance of seizure 336 freedom after the surgery and thus support the decision to proceed with surgery. On the contrary, a bad 337 ESSES prediction would indicate a 76.5% chance that the surgery would fail. This may be suggestive of 338 the need of more presurgical evaluations or a different resection strategy, and eventually indicate a low 339 probabily of complete seizure freedom after the surgery. 340

Finally, in order to test whether the information provided by the three biomarkers could be combined to 341 improve the prediction results, we performed a machine learning analysis using an adaptive boosting 342 algorithm with random undersampling and leave-one-out cross-validation (figure 7A,B). The input 343 variables for the classification algorithm were  $\delta IR(RA)$ ,  $S(R_{op})$  and  $Ov(RA, R_{op})$ . We found that, even 344 though the accuracy of the model was good (0.71) the machine learning model was biased towards the 345 majority class (SF), with only 35% of NSF cases correctly identified (precision = 0.37, sensitivity 346 = 0.35) and a poor result for F1 = 0.36, even though the considered algorithm (RUSboost) was designed 347 to correct for class imbalance. However, the minority class in our case contained only 8 cases, likely 348 preventing the model from being able to generalize. In order to address this issue, we created a *combined* 349 *cohort* (N = 49) pooling together the patients from the modeling and validation cohorts (figure 7C,D). 350 The combined cohort had 12 NSF cases (50% increase), and the new model was able to identify the 351 majority of NSF cases correctly (72% of SF cases and 63% of NSF cases). Even though the accuracy of 352 the model (0.70) did not improve, the remaining measures, which are less affected by class imbalance, 353 did (precision = 0.42, sensitivity = 0.63, F1 = 0.51). Overall, the machine learning model was not able 354 to improve upon the results found using the individual variables (see table 3), and indeed the prediction 355 was predominantly based only on one biomarker, namely the effect of the planned resection on the 356 modeled seizures,  $\delta IR(RA)$ . Due to the small sample size, we could not determine whether this was due 357 to intrinsic model limitations, suboptimal hyperparameters, or simply a too small group size (particularly 358 of the minority class). Our set-up (leave-one-out cross-validation combined with random undersampling) 359 was designed to minimize the effects of the small sample size, but could not avoid them fully. 360

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Prediction of surgical outcome: validation cohort. A Prediction results using each of the three model-based Figure 6. 361 biomarkers of surgical outcome: the size of optimal resections  $S(R_{op})$ , the overlap between optimal resections and the resection 362 area,  $Ov(R_{op}, RA)$ , and the decrease in seizure propagation due to simulation of the planned resection strategy,  $\delta IR(RA)$ . NSF 363 (SF) cases are shown by black (white) rectangles. The bottom row shows the fraction of biomarkers (0 - 3 out of 3) with a 364 positive (i.e. NSF) classification (refereed to as "Average model prediction" in the figure), for each patient. Surgical outcome 365 is shown in the top row. NSF cases are highlighted by a red arrow and by red labels. B Relative number of cases identified as 366 NSF by n biomarkers, n = 0, 1, 2, 3, respectively for the SF (blue, left-side bars, N = 26) and NSF (red, right-side bars, 367 N = 8) groups. 368

Metric	diff	rks	p
$\delta IR(R_{op})$	-4.34	411.5	0.03
$Ov(R_{op}, RA)$	0.20	495.5	0.03
$\delta IR(RA)$	0.26	513	0.02

374 Table 1. Summary of statistical comparisons: difference between SF and NSF groups (validation cohort). diff and rks stand respectively for the difference between the SF and NSF groups and the ranksum value. 375

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Figure 7. Prediction of surgical outcome using a machine learning algorithm (RUSBoost) with leave-one-out cross validation. As input variables we used the normalized decrease in seizure propagation after virtual resection of the RA,  $\delta IR(RA)$ , the size of optimal resections  $S(R_{op})$  and the overlap of optimal and clinical resections  $Ov(R_{op}, RA)$ . Panels A,B show the confusion matrix and predictor importance for the validation cohort (N = 34, 8 NSF), and panels C,D are for the combined cohort (N = 49, 12 NSF).

Variable	True negatives: SF		True	True positives: NSF		Prec.	Sensitivity	F1	AUC
$S(R_{op})$	19	0.73	6	0.75	0.74	0.46	0.75	0.57	0.71
$Ov(RA, R_{op})$	18	0.69	6	0.75	0.71	0.43	0.75	0.55	0.69
$\delta IR(RA)$	21	0.81	6	0.75	0.79	0.55	0.75	0.63	0.78

Table 2. Results of the classification analyses for the validation cohort. Results correspond to the optimal points of the ROC curves according to the Youden criterion to account for class imbalance. For each group (SF, NSF), we show the number of correctly identified cases by absolute number and relative frequency. The remaining columns correspond respectively to the accuracy (Acc.), precision (Prec.), sensitivity, F1 statistic and area under the curve (AUC).

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	Variable	True negatives: SF		True p	positives: NSF	Acc.	Prec.	Sensitivity	F1
	$S(R_{op})$	19/26	(= 0.73)	4/8	(= 0.50)	0.68	0.26	0.50	0.38
on	$Ov(RA, R_{op})$	18/26	(= 0.69)	5/8	(= 0.63)	0.68	0.38	0.63	0.51
lidati	$\delta IR(RA)$	21/26	(= 0.81)	5/8	(= 0.63)	0.76	0.50	0.63	0.57
Va	Combined	21/26	(= 0.81)	6/8	(= 0.75)	0.79	0.55	0.75	0.65
	RUSboost	(	).82		0.35	0.71	0.37	0.35	0.36
Combined	RUSboost	(	).72		0.63	0.70	0.42	0.63	0.51

Results of the prediction analyses for the validation and combined cohorts. For each analysis, we used a leave-Table 3. 380 one-out crossvalidation such that a predictive model was build to predict the outcome of each patient using the data from the 381 remaining N-1 patients. For the individual variables, the results correspond to the optimal points of the ROC curves according 382 to the Youden criterion. For the machine learning analyses, they were derived from an adaptive boosting (AdaBoost1, Matlab 383 2018) algorithm with leave-one-out crossvalidation, combined with random undersampling (RUSboost) to account for class 384 imbalance. Results were averaged over 10 iterations of the AdaBoost1 algorithm. For the combined method, the results from 385 the three individual analyses were combined, and a NSF classification was assigned to patients with at least two positive 386 (NSF) classifications. For each group (SF, NSF), we show the number of correctly identified cases by absolute number and 387 relative frequency. The remaining columns correspond respectively to the accuracy (Acc.), precision (Prec.), sensitivity and F1 388 statistic. For machine learning analyses only the average fraction of correctly predicted cases is shown in the true negatives and 389 true positives columns, since absolute results can vary per realization of the prediction algorithm. 390

#### DISCUSSION

Personalized models of brain dynamics can aid the treatment of patients with neurological disorders. In this study we presented ESSES (Epidemic Spreading Seizure and Epilepsy Surgery model): a framework to aid epilepsy surgery planning on a patient-by-patient basis. ESSES defines individualized seizure propagation models that integrate multimodal presurgical data, and can propose alternative resection strategies and provide confidence bounds for the probability of success of a given strategy. The implementation of ESSES in clinical practice may thus eventually improve the chances of achieving a good postsurgical outcome. <sup>398</sup> In this study we proposed a combined setting such that ESSES' parameters were fitted in a retrospective

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study (N = 15) using iEEG data of ictal activity, in analogy with previous studies (Goodfellow et al.

(2016); V. Jirsa et al. (2017); Kini et al. (2019); Makhalova et al. (2022); Moosavi et al. (2022);

H. E. Wang et al. (2023)). We validated that ESSES captured the main aspects of seizure propagation and 401 was able to reproduce the iEEG-recorded seizures, in agreement with our previous studies (Millán et al. 402 (2022, 2023)). Remarkably, the goodness-of-fit of ESSES-modeled seizures to iEEG data could identify 403 patients with a bad outcome with AUC = 0.79, 100% sensitivity and 57\% precision. Such information 404 may be integrated in the presurgical evaluation of the patients for whom iEEG data is available: different 405 resection strategies may be tested as the origin of the ESSES-modeled seizures (Millán et al. (2023)), 406 with a low goodness-of-fit being indicative of a low chance of seizure freedom. In particular, a bad 407 prediction by the model would indicate (in this cohort) a 57% chance of a bad outcome (to be compared 408 with only a 26.7% NSF rate in this cohort). Conversely, all patients identified as SF by the model could 409 proceed to surgery with high expectations (100% in this group) of good outcome. 410

The novel aspect of this study consisted of a subsequent *pseudo-prospective study* with an *independent* 411 cohort and in a blind setting. Importantly, we did not require the presence of iEEG data in the 412 pseudo-prospective study, and instead the multimodal presurgical information available for each patient 413 was integrated into seed-probability maps. In this manner ESSES can be adapted to the information 414 available for each patient, in a quantitative and systematic manner. IEEG data is highly invasive and 415 burdensome for the patient, and thus not always part of the presurgical evaluation. For instance, only 19 416 of the 34 patients of the validation cohort had undergone it. Thus, by not requiring iEEG data ESSES can 417 be applied to a much larger patient population than traditional approaches (Goodfellow et al. (2016); 418 V. Jirsa et al. (2017); Kini et al. (2019)), with the expected wider impact. 419

ESSES may be applied prospectively as follows. First of all, ESSES may suggest optimal resection
strategies, in analogy with previous studies (An et al. (2019); Laiou et al. (2019); Millán et al. (2022);
Nissen et al. (2021)), with the advantage that all multi-modal presurgical information available for each
patient is integrated into ESSES, instead of considering only one source used for network reconstruction.
We note that these resections are optimal within the framework of the model, and this does therefore not
guarantee optimal clinical outcome. Nevertheless, we have found that these virtual resections have good
predictive value of surgical outcome. The *optimal resection strategy*, defined here as the smallest

resection leading to a 90% decrease in (modeled) seizure propagation, can be used as a first indicator of the chances of seizure-freedom after *any* surgery. In our pseudo-prospective predictive framework (emulating the presurgical conditions) the size of this resection could predict 50% of patients with bad outcome (table 3), whereas the relative NSF rate in this group was 23.5%. This result is independent of the resection strategy and it is completely characterized by the presurgical information available for each patient. Thus, a bad prognosis could indicate that either the presurgical information available is not of sufficient quality, or that the patient is unlikely to be seizure-free with any resection strategy.

ESSES can also provide information about the prognosis after a particular resection by i) comparing it to 434 the optimal ESSES resection strategy and ii) quantifying its effect on seizure propagation in the 435 patient-specific ESSES model. Here we found that resections with a larger overlap with the optimal 436 virtual resection were more likely to lead to seizure freedom, in agreement with previous studies 437 (Goodfellow et al. (2016); Kini et al. (2019); Makhalova et al. (2022)). Similarly, resections leading to a 438 larger decrease in seizure propagation in ESSES were associated with a larger probability of 439 seizure-freedom after the resection, in agreement with other modeling (Goodfellow et al. (2016); Kini et 440 al. (2019)) and network-based (Bartolomei et al. (2017); Lopes et al. (2017); Nissen et al. (2017)) studies. 441 Here we considered only the planned resection strategy, which was approximated here by the resection 442 area, since this information could be derived in a systematic manner, and this set-up allowed us to 443 validate ESSES' findings. In a presurgical setting, different strategies could be tested to measure the 444 probability of seizure freedom after each one. In particular, we found that, when combining the 445 information from the three model-based biomarkers (namely the size of the optimal resection, its overlap 446 with the planned resections, and the effect of the planned resection on modeled seizure propagation) 447 could predict pseudo-prospectively 81% and 75% of SF and NSF cases (see table 3), whereas the relative 448 group ratios were 76.5% and 23.5%, respectively. Clinically, this implies that if a good prognosis is found 449 by at least two biomarkers, then there is a 91.3% (true negative rate, 21 cases were SF of the 23 predicted 450 by the model) chance that the patient will be seizure-free, and the patient can proceed with the surgery 451 with the knowledge that they will likely have a good outcome. Conversely, a bad prognosis by at least 452 two biomarkers indicates a 55% chance of bad outcome, and may be interpreted as an ESSES suggestion 453 to perform more presurgical testing or consider alternative resection strategies. Importantly, epilepsy 454 surgery may still improve the quality of life of the patient even when complete seizure freedom can not 455

<sup>456</sup> be achieved. Thus, moderate *a priori* chance of a bad outcome is not necessarily a contraindication for <sup>457</sup> surgery, but it is important in the presurgical counseling of the patients.

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Our findings here did not depend on the presence of iEEG data, and even when iEEG data were available 458 we only included a low-resolution description of them. IEEG data does provide the most detailed 459 information of epileptogenic activity, and is it often the most valuable tool to identify the epileptogenic 460 zone or predict surgical outcome for patients with complicated ethiology (Bernabei et al. (2023); 461 Gunnarsdottir et al. (2022); Makhalova et al. (2022); Proix et al. (2017); Runfola et al. (2023); Sinha et 462 al. (2017); Y. Wang et al. (2023)). In fact, for the modeling cohort we found the best classification results 463 when using the goodness-of-fit of ESSES-predicted seizure propagation patterns to the iEEG seizures, in 464 agreement with previous studies (Makhalova et al. (2022)). IEEG imaging however is burdensome to the 465 patient, has risk of complications, and has limited spatial coverage. A first prediction of surgical outcome 466 could thus be performed with ESSES when the results of non-invasive testing have been obtained, and an 467 iEEG study might be avoided if the model already predicts a good outcome with the existing data. 468

In summary, we showed here that ESSES could identify patients with good outcome presurgically based 469 on i) the smaller size of the optimal ESSES resection strategies, ii) a larger overlap of the planned 470 resection strategy with the optimal ESSES resection, and iii) a larger effect of the planned resection 471 strategy on decreasing (modeled) seizure propagation. Our findings here indicate that ESSES could be 472 generalized to other patient populations (as we did with the validation cohort), with the only requirement 473 of a patient-specific brain network, and can incorporate multimodal information from the existing 474 presurgical evaluation, in particular without requiring the presence of iEEG data. The ESSES-based 475 biomarkers identified here could be taken into account during presurgical planning to evaluate the need 476 for more testing, or may lead to the decision to forgo the surgery, if a bad outcome is predicted. This 477 extra information may be particularly valuable for patients with complicated ethiology (e.g. discordant 478 information from different modalities, variable seizure propagation patterns, multiple seizure onset 479 zones), for whom the discussion of whether or not to perform the surgery is challenging. 480

## **481** ESSES modeling framework

ESSES consists of different interconnected elements, namely i) the underlying network structure; ii) the
 seizure propagation model (and parameter fitting); iii) the seizure onset zone model; iv) the virtual

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resection model; and v) the virtual resection optimization algorithm. Each of these different elements was 484 designed to model a particular aspect of epilepsy surgery in a synergistic manner. For instance, the 485 emergent properties of the seizure propagation model (the SIR model) led the design of the virtual 486 resection optimization algorithm. At the same time, the modular organization of the framework allows 487 for the independent improvement or modification of each of the modules. In fact, different modules were 488 developed and analyzed in detail in our previous studies. For instance, the virtual resection algorithm 489 model was initially designed in Nissen et al. (2021) and improved in Millán et al. (2022), whereas the 490 seizure propagation and parameter fitting model as used here was mainly defined in Millán et al. (2023). 491 Below we discuss the main modeling considerations and results for each ESSES module. 492

As the underlying network structure we considered MEG-derived whole-brain networks as a proxy for 493 structural connectivity, following our previous works (Millán et al. (2022, 2023)), and in contrast with 494 other works (An et al. (2019); V. Jirsa et al. (2017); Nissen et al. (2021); Sip et al. (2021)). MEG provides 495 highly temporally resolved information with good spatial resolution and uniform coverage. Our previous 496 studies showed that MEG networks based on the amplitude envelope correlation (AEC) can integrate 497 information from both short-range structural connections (by not correcting for volume conduction) and 498 long-range functional coupling. Thus, AEC-MEG networks can be used as a cost-effective proxy for 499 structural connectivity (Millán et al. (2022)) with much lower computational cost than DWI 500 (Diffusion-Weighted Imaging) networks, whilst also being more sensitive to long range connections, in 501 particular inter-hemispheric ones, that may often be missed by DWI (Chen et al. (2015)). It would be an 502 interesting question for future studies to discriminate whether structural or functional connections drive 503 seizure propagation, in analogy to recent studies on the spreading of abnormal proteins associated with 504 Alzheimer's disease (Schoonhoven et al. (2023)). 505

The MEG networks were thresholded at different levels to prune out spurious connections, following previous studies (Millán et al. (2022, 2023); Nissen et al. (2021); Schoonhoven et al. (2023)). This requires the use of an arbitrary threshold, which we fitted to the iEEG data. In all cases we considered sparse networks (the maximum density considered was 0.35), and the operating point of ESSES was set at a very low density (0.03). This small density prevented weak or negative correlations from being included in the thresholded network. The proposed thresholding method can become a limitation if denser networks, including more connections, are considered. == D R A F T

ESSES was based on a simple epidemic spreading model, the SIR model. Epidemic spreading models, 513 such as the SIR or SIS (Susceptible-Infected-Susceptible) models, describe the basic aspects of spreading 514 phenomena on networked systems (Pastor-Satorras et al. (2015)), and have been used to describe other 515 neuro-physiological processes before, such as the spreading of pathological proteins on brain networks 516 (Peraza et al. (2019); Schoonhoven et al. (2023)) or the relation between brain structure and function 517 (Stam et al. (2016)). Epidemic spreading models have been extensively studied on different network 518 substrates (Pastor-Satorras et al. (2015)) and are supported by a well-grounded mathematical and 519 computational framework that we can use to our advantage in the context of epilepsy surgery. For 520 instance, from an epidemic spreading perspective, it is to be expected that hub removal plays a major role 521 in the decrease of seizure propagation, as found experimentally (Lopes et al. (2017); Nissen et al. 522 (2017)), with the spreading threshold heavily influenced by the existence of hubs (Pastor-Satorras et al. 523 (2015)). This theoretical background guided the design of an efficient virtual resection optimization 524 algorithm, such that the decrease in seizure propagation after a virtual resection could be approximated 525 by the decrease of centrality of the seed regions. 526

As we showed here and in previous works, epidemic spreading models can also reproduce the 527 fundamental aspects of seizure propagation at the whole-brain level in epilepsy patients (Millán et al. 528 (2022, 2023)). As ESSES's working point we chose here the values of the global parameters that led to 529 the maximum average goodness-of-fit of the modeling cohort (figure 1). Importantly, ESSES was still 530 individualized for each patient by means of the patient-specific brain connectivity, setting the local 531 spreading probabilities, and the patient-specific seed regions (based on the seed-probability maps built 532 with multi-modal presurgical information). As we showed in our previous study (Millán et al. (2023)) 533 and in the supplementary information here (Supp. section 5.2), by not individualizing the global model 534 parameters (namely  $\rho$  and  $\gamma$ ) for each patient we were able to reduce noise effects by integrating together 535 ictal data from different patients. Moreover, this formulation allowed us to generalize ESSES to patients 536 for whom iEEG seizure-propagation patterns were not available. 537

<sup>538</sup> Our findings in this study indicated that the iEEG seizure propagation patterns were significantly better <sup>539</sup> explained by ESSES for SF patients, and in fact all NSF cases could by identified by a bad ESSES fit, <sup>540</sup> and 73% of the SF cases by a good fit. There are several possible explanations for these findings. Given <sup>541</sup> that the epidemic seed was based on the resection area for each patient in this part of the analyses, a

simple explanation is that the resection strategy might have been better for SF patients given the existing 542 information. However, the difference could also arise from the iEEG data: the sampling may have been 543 inadequate for NSF patients (Sip et al. (2021)), or these may have presented seizure dynamotypes (Saggio 544 et al. (2020)) that were not well-explained by the considered epidemic spreading model (SIR model). The 545 fact that the optimization of virtual resections analysis –which did not depend on the clinical resection 546 area- also found differences between the SF and NSF groups points towards an intrinsic difference 547 between the presurgical data of the two groups, and not only to a sub-optimal surgical strategy for the 548 NSF group. 549

The next ingredient of ESSES was the definition of the seizure onset zone in the model, that is, the set of brain regions from which seizures originate. In this study we presented a method to combine the multimodal presurgical information available for each patient into *seed-probability maps*. This set-up thus emulated the clinical situation prior to the surgery, where a surgical strategy has been devised based on the information that is available from the presurgical evaluation. It would also allow for flexibility in the clinical application of ESSES: if more evaluations become available these could be readily integrated into the seed-probability map to update ESSES's results.

The final key ingredients of ESSES were the simulation and optimization of resection strategies. Here we 557 considered a node-based resection such that the resected nodes were disconnected from the network. This 558 approach however does not take into account possible widespread effects or plasticity mechanisms, 559 which could also be included into the model (Demuru et al. (2020)). The virtual resection optimization 560 algorithm was originally validated in our previous studies (Millán et al. (2022); Nissen et al. (2021)). 561 Given that optimizing virtual resections is highly computationally demanding, the algorithm took 562 advantage of the mathematical link between network structure and SIR dynamics to reduce the 563 dynamics-based optimization problem (i.e. finding the resection leading to a minimum seizure 564 propagation) into a network optimization problem (i.e. finding the resection leading to a minimum seed 565 efficiency). This was also motivated by our previous finding that the effect of a resection on the model 566 depended strongly on the centrality of the seed regions after the resection (Millán et al. (2022); Nissen et 567 al. (2021)). In particular, Nissen et al. (2021) found that removing connections to the network hubs was 568 the most efficient way to decrease seizure propagation, whereas Millán et al. (2022) verified a strong 569 correlation between a decrease in closeness centrality of the seed and a decrease in seizure propagation 570

<sup>571</sup> following a virtual resection. The effect of a resection on seizure propagation is also influenced by other
<sup>572</sup> network and model properties, and as a consequence the optimal network-based and SIR-based
<sup>573</sup> resections may differ slightly (Millán et al. (2022)). However, the intrinsic noise in the seed definition, in
<sup>574</sup> the seed-probability maps, and in the actual origin and propagation patterns of iEEG-recorded seizures
<sup>575</sup> created variability in the clinical data that absorbed the differences between the network-based and
<sup>576</sup> SIR-based optimal resections (which we previously found to be small anyway (Millán et al. (2022))).

The virtual resection optimization algorithm considered here imposed no conditions on the location of the resected regions, nor did it force that the resection strategy was made up of only one set of adjacent regions. Conditions on the resection strategies could be imposed, such as preserving eloquent cortex or forbidding bi-hemispheric resections (An et al. (2019); Laiou et al. (2019)). This would limit the dimensionality of the space of possible resection strategies and simplify the computations. However, by not imposing any conditions here we derived an *optimal* ESSES resection against which other, perhaps clinically more realistic, strategies could be tested (by e.g. measuring their overlap as we did here).

### 584 Modeling considerations and limitations

There are inherent limitations in the modeling of virtual resections, as the findings cannot be directly tested and we often rely on retrospective data. Here we have attempted to simulate how an epilepsy surgery model could be used in the clinic, i.e. prospectively, by considering only the presurgical information that is typically available to the clinical team. However, the optimal resections suggested by ESSES can still not be tested in practice, and in fact can only be considered optimal within the context of the model. Only long-term testing of the framework in the clinic can truly validate the use of computational models in epilepsy surgery.

ESSES is an abstraction of seizure dynamics that does not aim to reproduce the detailed
bio-physiological processes involved in seizure generation and propagation, but aims to focus only on the
most relevant features of seizure propagation (Millán et al. (2022, 2023); Nissen et al. (2021); Sip et al.
(2021)). In order to validate ESSES as a framework to simulate seizures, we compared the modeled
seizures with those recorded via iEEG. This required, however, a simplified representation of the iEEG
data. In particular, as there was no intrinsic time-scale in the SIR model, and to avoid introducing an
arbitrary one, we reduced the iEEG data to a pattern that describes the activation order of the sampled

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ROIs. Furthermore, even if ESSES provides a good representation of the iEEG seizures, extrapolating
these results to the simulation of the effect of a resection is not trivial. Moreover, our virtual resection
technique assumed that the effect of a surgery could be approximated simply by removing or
disconnecting the resected regions, whereas in practice widespread effects and compensation mechanisms
are expected (Demuru et al. (2020)). Here we validated ESSES' results against postsurgical outcome, but
seizure freedom is not a perfect gold standard either. For instance, in cases with a good outcome a smaller
resection could potentially also have led to seizure freedom (Millán et al. (2022); Nissen et al. (2021)).

All modeling frameworks are affected by the need to (sometimes arbitrarily) choose modeling 606 parameters, which go from the data reduction process to the choices of thresholds and metrics for the 607 final analyses. Here we considered well-established data preprocessing techniques (Hillebrand et al. 608 (2016)). ESSES was validated in previous studies (Millán et al. (2022, 2023); Nissen et al. (2021)), and 609 importantly we found that the results held for an independent cohort, and that modeling details (such as 610 the simulation algorithm for the SIR model) did not affect the main results (Millán et al. (2023)). A 611 simple model to simulate seizure propagation (the SIR model), also reduced the number of modeling 612 parameters so that the findings could be more easily generalized. Some arbitrary choices were still 613 needed, such as the definition of the 90% threshold to select the optimal resection strategy. However we 614 validated that similar results were obtained when another resection (the disconnecting resection) was 615 considered. 616

The seed-probability maps were based on an existing low-resolution database (*Castor Electronic Data Capture.* (n.d.)). Seed regions were consequently widespread over the network. This also led to a large variability in the results of different simulations for each patient (see for instance figures 4A,B and 5A), as these depended strongly on the seed realization. In order to improve the resolution of the model and minimize noise, the data from each modality could be integrated directly into the model, skipping the 34-region description in the database.

Finally, a limitation of this study is the small size of the non-seizure-free group, with only 4 cases in the modeling cohort and 8 in the validation cohort. This small size limited the classification and prediction analyses, and prevented us from building a more sophisticated machine learning model based on our analysis. With the proposed leave-one-out-crossvalidation method, combined with random undersampling and a small input space (only three data-points per patient), we attempted to overcome these limitations, but we were not able to improve upon the simpler ROC-based prediction results. Future studies involving more that one center have the potential to at least diminish this limitation.

# **CONCLUSION AND OUTLOOK**

Individualized computational models of seizure propagation and epilepsy surgery based on 630 patient-specific brain connectivity can reproduce individual iEEG seizure propagation patterns and aid 631 epilepsy surgery planning by proposing alternative resection strategies and providing estimates on the 632 likelihood of seizure freedom after the surgery. Here we presented the ESSES framework for seizure 633 propagation and epilepsy surgery. ESSES combines SIR epidemic spreading dynamics over 634 patient-specific MEG brain connectivity with a virtual resection framework. We defined a method to 635 derive patient-specific regional epileptogenicity maps from the presurgical evaluations of the patients in a 636 systematic and quantitative manner, and integrated them into ESSES. We performed a 637 pseudo-prospective study emulating the use of ESSES in clinical practice, prior to surgery. In the 638 pseudo-prospective analyses we did not require the presence of iEEG data, demonstrating that the model 639 could be applied to larger patient populations. We found that the goodness-of-fit of ESSES to the iEEG 640 seizures (in a retrospective study), the effect of the planned resection strategy, as well as the size of 641 ESSES optimal resections and their overlap with the planned resection, predicted surgical outcome with 642 0.68 - 0.76 AUC and 0.50 - 0.63 sensitivity to identify non-seizure-free patients. Our results thus 643 prescribe the use of ESSES during the presurgical evaluation to evaluate the need for further presurgical 644 testing on a case-by-case basis or, conversely, support the decision to proceed with surgery in the case of 645 a good-outcome prediction. For cases where a bad outcome is predicted, the surgical plan may be altered 646 to include ESSES's results. 647

## **METHODS**

The general design of the study is detailed in figure 2 and Supp. figure 7. Namely, we first set the hyperparameters of ESSES using a *modeling cohort* (N = 15) for which seizure propagation patterns derived from iEEG recordings were available. Then, ESSES was fitted with multimodal patient-specific data (in the form of seed-probability maps), and it was used to a) identify optimal resection strategies for each patient and b) predict the chance of a good outcome after a given resection. Then, ESSES was applied to a *validation cohort* (N = 34) in a pseudo-prospective analysis with a blind setting to emulate the presurgical conditions. That is, during the application of ESSES to determine optimal resection strategies, the researchers were blind to the actual clinical resection and surgical outcome of each patient. This data was subsequently de-blinded in two stages. First, the resection areas were obtained to be used as a proxy for the surgical plan of each patient to a) compare them with ESSES's optimal resection strategy, and b) simulate the effect of the surgical plan in ESSES. Finally, we de-blinded the one-year surgical outcome to enable a statistical validation of the results.

#### 660 Patient groups

We included two patient groups in this study, the *modeling cohort* for the model definition (retrospective study) and the *validation cohort* for the pseudo-prospective validation. All patients had undergone resective surgery for epilepsy at the Amsterdam University Medical Center, location VUmc, between 2013 and 2019. All patients had received an MEG recording, and underwent pre- and post-surgical magnetic resonance imaging (MRI). All patients gave written informed consent and the study was performed in accordance with the Declaration of Helsinki and approved by the VUmc Medical Ethics Committee. The excluding criterion was the existence of a prior brain surgery.

Both patient groups were heterogeneous with temporal and extratemporal resection locations and 668 different etiology (see Supp. tables 1 and 2 for details). Surgical outcome was classified according to the 669 Engel classification at least one year after the surgery (Engel Jr (1993)). Patients with Engel class 1A 670 were labelled as seizure-free (SF), and patients with any other class were labelled as non-seizure-free 671 (NSF). The modeling cohort consisted of 15 patients (4 NSF, 11 females) who had also undergone an 672 iEEG (invasive electroencephalography) study, including post-implantation CT-scans. This same cohort 673 was already included in Millán et al. (2023), and partially in Millán et al. (2022). The validation cohort 674 consisted of 34 patients (8 NSF, 13 females). No extra requirements (other than the presence of an MEG 675 recording of sufficient quality) were placed. In order to maintain the pseudo-prospective setting, the 676 research team was blind to the resection area and outcome of the validation cohort patients. In order to 677 perform the final analyses, for which this information was needed, the data was coded to avoid 678 identification. For two cases of the validation cohort (cases 2 and 9) the data of surgical outcome was 679

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de-blinded together with the data of the resection area as the research team became aware of a subsequent resective surgery (indicative of a bad outcome of the first surgery).

#### 682 Individualized Brain Networks

Seizure propagation was modeled on the patient-specific brain networks, as derived from MEG data, for 683 both cohorts (see Supp. figure 7). For each patient, a 10 to 15 minutes eyes-closed resting-state (supine 684 position) MEG recording was used to derive broadband (0.5 - 48.0 Hz) MEG functional connectivity. All 685 instrumental and methodological details were equal to our previous studies (Millán et al. (2022, 2023)) 686 and are detailed in the supplementary information (Supp. section 2). Functional networks were generated 687 considering each of the 246 ROIs of the Brainnetome (BNA) atlas (Fan et al. (2016)) as nodes. The 688 elements  $w_{ij}$  of the connectivity matrix, indicating the strength of the connection between ROIs i and j, 689 were estimated by the AEC (Amplitude Envelop Correlation) (Brookes et al. (2011); Bruns, Eckhorn, 690 Jokeit, and Ebner (2000); Colclough et al. (2016); Hipp, Hawellek, Corbetta, Siegel, and Engel (2012)), 691 without including a correction for volume conduction. The uncorrected AEC maintains information 692 about the structural connections, which are mainly determined by the distance between each ROI pair, by 693 not correcting for volume conduction. We validated the relationship between AEC-MEG and structural 694 networks in a previous study (Millán et al. (2022)) by comparing them with a well-validated model for 695 structural connectivity: the exponential distance rule (EDR) network. Based on animal studies, the EDR 696 specifies that the weights of structural connections in the brain,  $w_{ij}$ , decay exponentially with the 697 distance between the ROIs  $d_{ij}$  (Ercsey-Ravasz et al. (2013); Gămănuţ et al. (2018); Theodoni et al. 698 (2022)), i.e.  $w_{ij} \propto exp(-\alpha d_{ij})$ . Recent studies have corroborated this behavior also in human structural 699 connectivity (Deco and Kringelbach (2020); Deco et al. (2021); Roberts, Perry, Roberts, Mitchell, and 700 Breakspear (2017)), although the EDR cannot capture all details of white matter connectivity, as this is 701 not isotropic (Betzel and Bassett (2018); Jbabdi, Sotiropoulos, Haber, Van Essen, and Behrens (2015); 702 Markov et al. (2013)), and includes long-range connections that are missed by the EDR (Roberts et al. 703 (2016)). However, the EDR is enough to capture the overall scaling of structural connections with the 704 distance as observed in the human structural connectome. In Millán et al. (2022) we validated that 705 AEC-MEG networks were strongly correlated ( $R^2 = 0.50$ ) with the corresponding EDR networks, 706 therefore showing that AEC-MEG reproduces at least partially the overall organization of structural 707 connectivity. Moreover, AEC-MEG networks also include long-range connections that may promote 708

<sup>709</sup> seizure propagation, but that may be missing from structural (i.e. DWI) networks (Jones, Knösche, and
 <sup>710</sup> Turner (2013); Reveley et al. (2015)). Thus, uncorrected AEC-MEG networks are a convenient way to
 <sup>711</sup> construct a network that resembles a structural network and includes long-range connections.

AEC values were re-scaled between 0 (perfect anti-correlation) and 1 (perfect correlation), with 0.5712 indicating no coupling (Briels et al. (2020)). Functional networks were thresholded at different network 713 densities  $\rho$  indicating the fraction of links remaining in the network. We note that the networks were 714 thresholded but not binearized, so that  $w_{ij}$  could take values between 0 and 1. The density thresholds 715 were chosen to be logarithmically distributed between 0.01 to 0.35. The weakest non-zero link included 716 in the network had an average weight of 0.54 (range: 0.52 - 0.56) for  $\rho = 0.35$ . At ESSES's operating 717 point (best model fit) the density was  $\rho = 0.03$ , and the weakest non-zero weight was 0.71 (range: 0.67 -718 0.76). 719

## 720 Resection Area

The resection area (RA) was determined from the three-month post-operative MRI. For the modeling cohort the resection areas were obtained as part of two previous studies (Millán et al. (2022, 2023)). For the validation cohort, to maintain a completely blind setting for the first analysis (*Optimization of alternative resections*), the resection areas were obtained during a second pre-processing step, as described in figure 2. Cases 9 and 20 of the validation cohort underwent the post-operative MRI on a different MRI scanner at their resection center, respectively one day and three weeks after the surgery. Case 9 also lacked a 3-month postoperative MRI, an MRI from 2 years after the surgery was used instead.

The post-resection MRIs were co-registered to the pre-operative MRI using FSP FLIRT (version 4.1.6) 12 parameter affine transformation. The resection area was then visually identified and assigned to the corresponding BNA ROIs, namely those for which the centroid had been removed during surgery.

# 731 *iEEG Seizure Propagation Pattern*

Patients in the modeling cohort underwent invasive EEG recordings using stereotactic electrode
 implantation as described in Millán et al. (2023). One characteristic iEEG-recorded seizure from each
 patient was used to derive a seizure propagation pattern in terms on the BNA ROIs, the *iEEG seizure pattern*, as described in Millán et al. (2023) and in the Supp. section 3.

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Authors: Millán et al.

#### 736 Seizure Propagation Model

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ESSES was based on our previous studies (Millán et al. (2022, 2023); Nissen et al. (2021)) where we 737 showed that simple epidemic spreading models could reproduce the spatio-temporal seizure-propagation 738 patterns derived from invasive EEG recordings, and that they could be used to simulate the effect of 739 different resection strategies in silico. ESSES was based on a well-known epidemic spreading model: the 740 Susceptible-Infected-Recovered (SIR) model (Pastor-Satorras et al. (2015)), which was simulated on the 741 patient-specific MEG brain network. The SIR model simulated the propagation of ictal activity from a set 742 of seed regions that were set to be infected at the beginning of the simulation to the remaining nodes in 743 the network, and the subsequent recovery of infected nodes. The SIR dynamics were defined by two 744 parameters: the probability  $\beta_{ij}$  that each infected node *i* propagates the infection to a neighbour *j* 745  $(S \to I)$ , and the probability  $\gamma_i$  that each infected node *i* recovers  $(I \to R)$ . For simplicity, we considered 746 here a global recovery probability  $\gamma_i = \gamma$ , and spreading probabilities given by the MEG network 747 connectivity:  $\beta_{ij} = w_{ij}$ . Thus, the spreading rate was determined by the density of connections in 748 network  $\rho$ . The two control parameters of ESSES are thus the network density  $\rho$ , and the recovery 749 probability  $\gamma$ . Depending on the network structure, the epidemics can show different spatio-temporal 750 spreading profiles described by the probability  $p_i(t)$  that each ROI i becomes infected at step t. 751

 $\rho$  and  $\gamma$  were fitted to the iEEG seizure-propagation patterns at the group level. The resection area was set 752 as the seed of epidemic spreading, and an ESSES seizure propagation pattern was built that described the 753 set of infected and non-infected ROIs during the SIR-simulated seizures, as well as the order in which 754 infected ROIs became infected. In order to take into account the stochastic nature of the SIR dynamics, 755 the participation of each ROI was weighted by the fraction of realizations in which it was involved in the 756 simulated seizure (since different ROIs became infected in different realizations). The goodness-of-fit of 757 the model,  $C(\rho, \gamma)$  (Millán et al. (2023)), quantified how similar the ESSES and iEEG patterns were. It 758 took into account two factors: the weighted correlation between activation orders of ROIs that were active 759 in both patterns,  $C_w$ , and the overlap between the active and inactive ROI sets of both patterns,  $P_{\text{overlap}}$ , i.e. 760

$$C = C_w \cdot P_{\text{overlap}}.\tag{1}$$

The details of this definition can be found in Supp. section 5.2.

We estimated C for a range of values  $\rho$  and  $\gamma$  logarithmically distributed (between 0.01 and 0.35 for  $\rho$ 762 and between 0.01 and 1.00 for  $\gamma$ ), considering  $N_R = 10^4$  iterations of the SIR dynamics 10 times in order 763 to determine average C values and their fluctuation for each patient. We then found the parameter set that 764 maximized C for each patient (see Supp. section 5.2 and Supp. figure 2) and at the group level (figure 765 1A). The model parameters that lead to the best fit at the population level defined the ESSES model and 766 were carried over to the pseudo-prospective analyses. Importantly, even though the SIR global 767 parameters were set equal for all patients, ESSES was individualized for each patient by means of their 768 patient-specific MEG brain connectivity, which defined the spreading probabilities, and their 769 patient-specific seed-probability map, which defined the seed regions. 770

The SIR dynamics was simulated by an adaptive Monte Carlo method (the BKL algorithm) in Matlab in discrete time, such that at each time step one new node became infected.  $N_R = 10^4$  iterations of the dynamics were run for each model configuration in all analyses.

### 774 Presurgical hypothesis of the seed regions

<sup>775</sup> We built seed-probability maps indicating the probability that each ROI started a seizure, for each patient
<sup>776</sup> of both cohorts. This is a key difference with our previous studies, where the seed regions were either
<sup>777</sup> derived from the resection area (Millán et al. (2022, 2023); Nissen et al. (2021)), which can only be
<sup>778</sup> known after the surgery, or from the iEEG data (Millán et al. (2022, 2023)). Here we defined a
<sup>779</sup> framework to integrate data from the different presurgical evaluations that were available for each patient,
<sup>780</sup> which was encoded in an existing database (Castor EDC, Ciwit B.V., Amsterdam (*Castor Electronic*<sup>781</sup> *Data Capture*, (n.d.))).

To compute the seed-probability maps, we considered the information available from 6 presurgical modalities: i) presence of ictal activity in EEG, ii) MRI lesions, iii) MEG abnormalities, iv) PET lesions, v) SPECT abnormalities and vi) iEEG recordings of ictal activity. All patients had undergone an EEG, MRI and MEG study, but not all of them presented PET, SPECT or iEEG data. The presence (1) or absence (0) of data of each modality was encoded in a variable  $D_m = 0, 1, m = 1, 2, ..., 6$ , for each patient.

<sup>788</sup> The database included information at the level of 34 regions, consisting of 6 frontal regions

789 (fronto-orbital, frontal-basal, frontal-parasagitaal, frontal-periventricular, frontal-lateral,

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<sup>790</sup> frontal-operculum), 6 temporal regions (hippocampus, amygdala, uncus, anterior-neocortical,

<sup>791</sup> posterior-neocortical, gyrus-parahippocampalis), 2 insular regions (anterior and posterior insula), 1

<sup>792</sup> central, 1 parietal and 1 occipital region, for each hemisphere. The temporal and frontal lobes are the

<sup>793</sup> most often involved in EZ and resection strategies, and thus are described in more detail in the database.

For each region i and modality m, the database indicates the presence (1) or absence (0) of abnormalities, 794 from which we derived binary abnormality maps  $a_{i,m} = 0, 1$ . The overall abnormality map  $A_i$  was 795 obtained by aggregating over all modalities available for each patient. Not all modalities are equally 796 relevant to establish the probability that a region is involved in epileptogenic activity: EEG is the least 797 focal, whereas iEEG provides the most localized information, and its results also integrate information 798 from the other modalities (as these affect where the iEEG electrodes are placed). In order to gauge these 799 differences, we weighted each modality m by a relevance factor  $\omega_m$ , with  $\omega = 1$  for EEG, 2 for MRI, 800 MEG, PET and ISPECT, and 4 for iEEG. Thus, the overall abnormality map was defined as 801

$$A_{i} = n^{-1} \sum_{m=1}^{6} D_{m} \omega_{m} a_{i,m},$$
(2)

where the normalization factor n is defined as  $n = \sum_{m=1}^{6} D_m \omega_m$ 

1

A clinician (ECWvS) defined a unique projection of the regions in the database on to the BNA ROIs. In 803 most cases the database regions corresponded to well-defined gyri that are also well-described in the 804 BNA documentation. A table describing the projection is included as supplementary material. We 805 projected the abnormality map  $A_i$  from the low-resolution description into the BNA atlas to obtain the 806 seed-probability maps  $SP_i$ , with i = 1, 2, ..., 246. Given that the description provided by the database 807 was broad and homogeneous (i.e. the considered ROIs are much larger than the BNA ROIs), and that 808 co-occurrence of abnormalities in different modalities is a strong indicator of the epileptogenic zone, we 809 included a re-scaling factor R to produce more focal seed-probability maps:  $SP_i = (A_i)^R$ , where j is the 810 region in the database corresponding to the BNA ROI *i*. We found that for R > 2 the results did not 811 depend strongly on R, and report here for R = 3. 812

#### **813** Virtual Resections

We conducted virtual resections of sets of nodes R by disconnecting them from the network, by setting to 0 all their connections. The effect of each resection was characterized by the normalized decrease in seizure propagation  $\delta IR(R)$  in the resected network (R) with respect to the original (0) one:

$$\delta IR(RA) = (IR_0 - IR_R)/IR_0,\tag{3}$$

where IR is the fraction of nodes that became infected at any point during the modeled seizure, namely,

$$IR = I(t \to \infty) + R(t \to \infty). \tag{4}$$

That is, IR takes into account all nodes that became infected during the simulated seizure, regardless of whether they eventually recovered or not. This characterizes the total extent of the simulated seizure.

- We performed two virtual resection studies, as detailed in figure 2. Firstly, we performed an *Optimization*
- of alternative resections analysis. We derived optimal virtual resections R of increasing sizes S(R)

(defined as the number of resected nodes) with an optimization algorithm based on simulated annealing 822 (Kirkpatrick, Gelatt, and Vecchi (1983)) and derived in our previous studies (Millán et al. (2022); Nissen 823 et al. (2021)). The optimization method took advantage of the relationship between SIR spreading and 824 network structure to use a structural metric -the seed efficiency- as a proxy for the actual effect of the 825 resection on seizure propagation  $\delta IR(R)$ . Thus, for each resection size S(R), the simulated annealing 826 algorithm searched for the resection R that minimized the seed efficiency  $E_R$ (seed) (Barrat et al. (2008); 827 Brockmann and Helbing (2013); Pinto, Thiran, and Vetterli (2012)).  $E_R$  (seed) measures the inverse 828 average distance from the seed nodes to the remaining nodes in the network: 829

$$E_R(\text{seed}) = \frac{1}{N_{\text{seed}}N_2} \sum_{i \in \text{seed}} \sum_{j \in \mathcal{S}_2} \frac{1}{d_{ij}},\tag{5}$$

where  $d_{ij}$  is the distance (in the network sense) between nodes *i* and *j*,  $S_2$  is the set of nodes that do not belong to the seed,  $N_2$  the size of this set, and  $N_{seed}$  the number of nodes that belong to the seed. In case of network disconnection, only nodes in the giant component were included in the seed and  $S_2$  sets.

All nodes were considered as possible targets of the resection. To compare between different patients we defined the normalized seed efficiency

$$e_R(\text{seed}) = E_R(\text{seed})/E_0(\text{seed}),$$
 (6)

where  $E_0$  (seed) is the seed efficiency in the original (un-resected) network. The actual effect of each resection was quantified by the seizure propagation level after the resection, IR(R), and the normalized 1

decrease in seizure propagation  $\delta IR(R)$ . We defined the *optimal ESSES resection*  $R_{op}$ , as the smallest resection leading to (at least) a 90% decrease in (modeled) seizure propagation. This resection was characterized by its size  $S(R_{op})$  and overlap with the resection area  $Ov(RA, R_{op})$ . We also defined the *disconnecting resection*  $R_D$  as the smallest resection that lead to seed disconnection (see Supp. section 5.4 and Supp. figure 6).

In the second virtual resection study, we simulated the effect of the planned resection for each patient, to measure its effectiveness in reducing seizure propagation. The resection area was used as a proxy for the resection strategy (figure 2: Simulation of the resection plan), since it could be derived in a systematic manner from the data.

For all virtual resection analyses the seed regions were derived from the patient-specific seed-probability 846 maps, and the underlying network was given from the patient-specific MEG network as before. In order 847 to obtain precise results, the effect of each resection was averaged over 300 independent realizations of 848 the seed regions from the seed-probability maps. As described in figure 2, for the validation cohort we 849 first performed the *Optimization of alternative resections* in a blind setting. Then the resection areas were 850 de-blinded and used as a proxy of the planed resection strategy to i) quantify the overlap of ESSES's 851 optimal resections with the resection strategy and ii) measure the effect of the planed resection in 852 decreasing (modeled) seizure propagation. Finally the one-year postoperative outcome was also 853 de-blinded and used for the statistical analyses. 854

## 855 Statistics

The weighted correlation coefficient was used to determine the correlation between the iEEG and ESSES seizure propagation patterns for the modeling cohort. In all analyses, for comparisons between SF and NSF patients, we used a two-sided Wilcoxon ranksum test. Significance thresholds for statistical comparisons were set at p < 0.05.

We performed receiver-operating characteristic (ROC) curve analyses to study the patient classification based on i) the goodness-of-fit of the model (modeling cohort), ii) the size of optimal and disconnecting resections (modeling and validation cohorts), iii) the overlap between optimal resections and the planed resection (modeling and validation cohorts), and iv) the effect of the planed resection on modeled seizure == D R A F T

<sup>864</sup> propagation (modeling and validation cohorts). A positive result was defined as bad outcome
 <sup>865</sup> (non-seizure-free, NSF) classification.

In order to account for the noise in the SIR model, the spreading dynamics were averaged over  $10^4$ 866 iterations of the SIR dynamics to derive each ESSES seizure pattern. The model fit analyses were 867 repeated 10 times to obtain averaged values. For the Virtual resection analyses we performed 300 868 independent realizations of the seed regions and SIR dynamics. Each seed realization was used to 869 measure seizure propagation in the original (before any resections) network and after the selected 870 resection of each size. For the Optimization of resections analysis we also ran the simulated annealing 871 algorithm 10 times for each resection size and selected the iteration that led to the minimal seed 872 efficiency. 873

For the classification analyses we report the accuracy= (TP + TN)/(TP + FP + FN + TN), precision TP/(TP + FP), sensitivity= TP/(TP + FN), F1 statistic (harmonic mean between precision and sensitivity) = 2TP/(2TP + FP + FN), and area under the curve AUC. For the prediction analyses, we built a predictive model for each patient using the data from the remaining patients, in a leave-one-out crossvalidation-type setting. The predictive model compounded the prediction results from these N = 34models. We measured its accuracy, precision, sensitivity and F1 statistic.

In the final analysis of the study we performed a predictive Machine Learning analysis based on the
AdaboostM1 algorithm (Matlab 2018) combined with random undersampling. AdaBoost is an adaptive
boosting machine learning algorithm in which the weights of mis-classified instances are adjusted
iteratively to improve the model. By combining adaptive boosting with random undersampling of the
majority class (SF group), the classification algorithm effectively addresses class imbalance and reduces
bias to the majority class and overfitting risks (*AdaboostM1 - Matlab 2018*. (n.d.); Friedman, Hastie, and
Tibshirani (2000)).

For each patient, three variables were considered as input for the prediction analysis: the size of the optimal resection  $S(R_{op})$ , its overlap with the resection area  $Ov(R_{op}, RA)$ , and the effect of the resection strategy on modeled seizure propagation  $\delta IR(RA)$ ). The goal of the machine learning algorithm was to predict surgical outcome. Due to the small cohort size, we performed a leave-one-out-cross-validation procedure, such that  $N_{pat}$  different training sets were created, each leaving out one patient, which was then used to test the prediction model. The training sets were formed by randomly undersampling the majority class (SF) to the size of the minority (NSF) class. The small cohort size also prevented us from including a validation set and performing parameter-tuning. Thus, we used default hyperparameters of AdaboostM1 (see *AdaboostM1 - Matlab 2018*. (n.d.) for details): the number of learners in each model was set equal to the group size minus one, the learning rate was set to 1.0 (default) and results were averaged over 10 iterations of the undersampling and AdaboostM1 procedures for each classification model. The machine learning analysis was performed twice: first considering only the patients in the validation cohort ( $N_{pat} = 34$ ), and secondly considering all patients (combined cohort,  $N_{pat} = 49$ ).

#### 900 Data availability

The data used for this manuscript are not publicly available because the patients did not consent for the sharing of their clinically obtained data. Requests to access to the data-sets should be directed to the corresponding author. All user-developed codes are publicly available on Github

904 https://github.com/anapmillan/ESSES.

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# **COMPETING INTERESTS**

<sup>915</sup> The authors declare that they have no competing interests.

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